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Medicare Non-Payment of Hospital-Acquired Infections: Infection Rates Three Years Post Implementation

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Background: Medicare ceased payment for some hospital-acquired infections beginning October 1, 2008, following provisions in the Medicare Modernization Act of 2003 and the Deficit Reduction Act of 2005.

Objective: We examined the association of this policy with declines in rates of vascular catheter-associated infections (VCAI) and catheter-associated urinary tract infection (CAUTI).

Data: Discharge data from the Florida Agency for Healthcare Administration from 2007 to 2011.

Study Design: We compared rates of hospitalacquired vascular catheter-associated infections (HA-VCAI) and catheter-associated urinary tract infections (HA-CAUTI) before and after implementation of the new policy (January 2007 to September 2008 vs. October 2008 to September 2011). This pre-post, retrospective, interrupted time series study was further analyzed with a generalized hierarchical logistic regression, by estimating the probability of a patient acquiring these infections in the hospital, post-policy compared to pre-policy.

Principal Findings: Pre-policy, 0.12% of admitted patients were diagnosed with CAUTI; of these, 32% were HA-CAUTI. Similarly, 0.24% of admissions were diagnosed as VCAI; of these, 60% were HA-VCAI. Post-policy, 0.16% of admissions were CAUTIs; of these, 31% were HA-CAUTI. Similarly, 0.3% of admissions were VCAIs and, of these, 45% were HA-VCAI. There was a statistically significant decrease in HA-VCAIs (OR: 0.571 (p < 0.0001)) post-policy, but the reduction in HA-CAUTI (OR: 0.968 (p < 0.4484)) was not statistically significant. Conclusions: The results suggest Medicare nonpayment policy is associated with both a decline in the rate of hospital-acquired VCAI (HA-VCAI) per quarter, and the probability of acquiring HA-VCAI post- policy. The strength of the association could be overestimated, because of concurrent ongoing infection control interventions.

Keywords: Medicare, Hospital-Acquired Infections, Patient Outcomes

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Introduction

Changes in Medicare policies have affected the health care system in a variety of ways. Earlier reforms focused on payments, with attention recently shifting to quality of care. Medicare used pay-forperformance in the Premier Demonstration Project to encourage hospital reporting of process measures as a means of improving patient safety (Petersen *et al.*, 2006). The Premier Demonstration Project was part of a value purchasing program initiated by Congress (Centers for Medicare & Medicaid Services, 2008), in response to reports by the Institute of Medicine about the preventability of some of the estimated 1.7 million U.S. hospital-acquired infections in 2002, and the resulting 100,000 deaths (Institute for Health Metrics, 2009; Scott, 2002).

As part of the value purchasing program, Medicare implemented a policy on October 1, 2008 that penalizes hospitals if Medicare patients acquire any of 8 conditions during their inpatient stay; i.e., if these infections were not present on admission. These conditions include catheter-associated urinary tract infections and vascular catheter-associated infections, two of three hospital-acquired infections that are either expensive to treat or commonly acquired. Under the new policy, hospitals would not be able to classify these conditions within a higher diagnosis related group (DRG) in order to receive additional reimbursement from Medicare (Institute for Health Metrics, 2009; Centers for Medicare & Medicaid Services, 2006). The Centers for Medicare & Medicaid Services (CMS) hopes that this policy will encourage hospitals to provide better quality of care by minimizing the rates of adverse events (Stone et al., 2010).

Researchers are divided over the likely impact of this policy on patient outcomes. McNutt *et al.*

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(2010), predict that the policy is likely to have a substantial impact on hospitals' financial performance and hence impact their behavior. The study used administrative hospital discharge data from all patients in 86 academic medical centers across the U.S. between July 2005 and June 2007, to estimate the average amount of reimbursement loss per hospital based on Medicare non-payment policy requirements. Using simulations, the study found that about 4% of patients had at least one of the targeted eight hospital-acquired conditions, and estimated the average amount of reimbursement loss for the 86 hospitals at approximately \$50 million per year (an average loss per case ranging from \$1,548 to \$7,310).

McNair *et al.* (2009), on the other hand, suggested a lower impact, because of a lower predicted financial loss per hospital. The study based estimates on a simulation model using the California Office of State Health Planning and Development 2006 Patient Discharge Dataset. They estimated that the Medicare non-payment policy would result in a reimbursement loss in California hospitals each year between \$92,000 and \$227,000 (resulting in about \$1.1 to 2.7 million nationally), substantially lower than what McNutt *et al.* (2010) estimated. Meddings *et al.* (2010) agreed with McNair that the expected impact will be minimal, due mainly to medical claims coding issues.

Conceptually, the policy is structured to provide an incentive for hospitals to improve their infection control practices in order to avoid financial loss. Predictors of hospital-acquired infections include patient characteristics; for example, sex, age, co-morbidities like renal disease and immune-compromised conditions (Kaye *et al*, 2011; Greene *et al.*, 2012), and hospital factors; for example, the type and duration of catheters left on a patient (Apisarnthanarak *et al.*, 2007; Johnson, Kuskowski, & Wilt, 2006). These studies and others have shown that hospitals can reduce infection rates when they follow certain proven guidelines. For example, in the keystone project, following guidelines like hand hygiene, removing unnecessary catheters, and cleaning the skin with chlorhexidine led to a significant reduction in central-line blood infections, up to zero infections (Pronovost et al., 2006). Active surveillance and effective control measures were also found to be associated with lower infection rates (Haley et al., 1980). Following these guidelines, however, may involve financial investments that are more costly than the predicted financial loss from the policy. Additional financial investments in infection control will, in turn, depend on the perceived financial strain of the policy on financial performance. In a survey of hospital CEOs regarding this policy, 70% planned to make financial investments by introducing new technologies to help improve infection rates (Pacemakerclub, 2012). Hospitals may improve outcomes due to the increased attention to patient safety brought about by the policy, and/ or from the perceived financial losses due to the policy. There are concerns, however, that the policy could result in unintended consequences, such as patient selection bias and overuse of antibiotics, which leads to antibiotic resistance (Burke, 2003).

Although researchers have looked at the likely financial impact of the policy and unintended consequences, few have reported actual outcome data in relation to the policy's implementation. The goal of this study is to evaluate, using both pre- and post-policy data, the association between Medicare's non-payment policy and patient outcomes for hospital-acquired catheter-associated urinary tract infections (HA-CAUTI) and vascular catheter-associated infections (HA-VCAI). This study will contribute to our understanding of the possible effect of Medicare's negative financial incentive on infection rates.

Methods

This study uses trend analysis (interrupted time series) and a pre-post design with retrospective data to examine rates of HA-CAUTI and HA-VCAI per quarter, and before and after implementation of the Medicare policy in October, 2008 (Wagner, Soumarai, Zhang, & Ross-Degnan, (2002). We hypothesized that both the rates and the individual risk of HA-CAUTI and VCAI will be lower after implementation of the policy.

Data

The primary source of data was administrative hospital discharges maintained by the Florida Agency for Healthcare Administration (Agency for Health Care Administration, 2010) from 2007 to 2011. Florida is an ideal state to consider as it was one of the few states (others include California and New York) that required their hospitals to report whether infections were present on admission (POA) before the Medicare requirement (Wald & Kramer, 2007). We did not include data prior to 2006, because the POA indicator and VCAI coding became a requirement in 2007; the VCAI coding was reported from the fourth quarter of 2007. Therefore data used was from quarter 1, 2007 for CAUTI and quarter 4, 2007 for VCAI.

Since 1988 AHCA has maintained a hospital discharge dataset with data on approximately two million patients per year. For this study, the dataset included 2,563,643 discharges in 2007, 2,571,688 in 2008, 2,491,033 in 2009, 2,604,090 in 2010, and 1,994,466 in the first three quarters of 2011. The dataset provided information on patient demographics, some hospital characteristics, and diagnosis codes. The study included all acute care hospitals, but excluded hospitals exempt from the policy, such as critical care hospitals, psychiatric

hospitals, long-term care hospitals, and special hospitals like children and cancer care hospitals (McHugh, Martin, Orwat, & Van Dyke, 2011).

Variables

The first objective of the study is to compare the discharged diagnosis rates of the targeted hospitalacquired infections per quarter, and before and after Medicare implemented the policy. Secondly, the study estimates the odds of an individual patient acquiring these infections in the hospital post-policy, while controlling for other patient characteristics.

Dependent Variables

Among the eight conditions targeted by the policy are CAUTI and VCAI. Medicare specified ICD-9 codes and procedure codes necessary to identify these conditions, as well as a POA indicator to determine whether the condition was hospitalacquired, which is defined as a condition not present upon admission, but that presents after 48 hours in the hospital (ECRI Institute Special Advisory, 2008).

For each discharge in the dataset, the presence or absence of CAUTI and VCAI were defined by ICD9 codes 996.64 and 999.31, respectively, and procedure codes 57.94 and 97.95. Then, each CAUTI and VCAI were further categorized as hospital-acquired or not based on the POA indicator (yes vs. no, or unknown).

Independent variables

The main independent variable, pre-or postpolicy, was determined for each discharge based on whether the discharge date was between January 1, 2007 and September 30, 2008 (pre-policy) or between October 1, 2008 and September 30, 2011 (post-policy). Other covariates included the individual patient's age (continuous variable), race (White, Black, or Other), gender (male or female), length of stay (continuous variable, we included only observations within 3 standard deviations of the mean), insurance status of the patient (Medicare, Medicaid, Private, and other Government insurance like Tri-Care), and whether the patient required ICU care (yes or no). The specific hospital for each discharge was also identified. We tried other models that included hospital characteristics (e.g., ownership, size, nurse full-time equivalent (FTE) and the number of co-morbidities per patient, but they made no difference on the results.

Analyses

The analysis first determined the average quarterly discharge rates of any diagnosed CAUTI and VCAI as defined by the policy, and the proportion of all HA-CAUTI and HA-VCAI. The average rates of the proportion of HA-CAUTI/CAUTI and HA-VCAI/VCAI were compared pre- and postpolicy implementation using an interrupted time series of the quarterly rates. The segmented regression analysis of interrupted time series, as described by Wager *et al.*, is of the following form:

 $\begin{aligned} Y_t &= \beta_0 + \beta_1 \times \text{Time} + \beta_2 \times \text{Intervention} + \beta_3 \times \\ \text{Time_after_Intervention} + e_t. \end{aligned}$

- Y_t is the outcome of interest (proportion of HA-CAUTI/CAUTI or HA-VCAI/VCAI per Time [in quarters]);
- β_1 , represents the base trend;
- Intervention is a dummy (0, 1) to distinguish between pre-and post-policy eras;
- β₂ estimates the immediate change (change in level), due to the intervention;
- Time_after_Intervention is 0 in the pre-policy era, and a continuous counting variable starting from 1 in the post-policy era;
- β₃ estimates the change in trend in the postpolicy era;

- the intercept (β_0) estimates the base level of the outcomes at the beginning of the series; and
- e_t is the error term.

Secondly, the analysis estimated the individual odds of being diagnosed with these infections (HA-CAUTI and HA-VCAI) post-policy, using a multivariate logistic regression in a generalized hierarchical linear model (GHLM).

GHLM approach gives more robust results, because it factors in the clustering effect at the hospital level (Raudenbush & Bryk, 2002).

The level-one covariates included individual patient race, age, type of insurance, gender, length of stay, and whether the patient required the intensive care unit. The level-two (cluster) variable was the specific hospital.

Results

Descriptive statistics

For the sample used in the CAUTI model, mean patient age was about 71 years, 78% were White, and 60% were male. Medicare (including about 8% Medicare advantage admissions) was the most common primary source of payment (81%), with only 8% having private insurance. About 38% of the patients used the ICU and had an average length of stay of 10 days.

The VCAI patients were younger on average (about 55 years old), a lower proportion had Medicare as the primary source of payment (about 52%–55%), and a much higher number compared to CAUTI patients had private insurance (about 21%–24%). Length of stay for VCAI patients was about 22 days and more than 50% used the ICU, indicating a much sicker population (Exhibit 1). There were 204 and 210 hospitals with CAUTI and VCAI claims respectively, pre-policy, but there were 225 hospitals with CAUTI and VCAI claims with CAUTI and VCAI claims post-policy.

	CAUTI ^a		VCAI ^b					
	Pre-	policy	Post-j	policy	Pre-p	policy	Post-j	policy
Category Description	N =	5,510°	N = 1	2,688°	$\mathbf{N} = 0$	5,119°	N = 2	3,703°
Percentage of admissions with infection	0.1	12%	0.1	6%	0.1	4%	0.3	0%
Proportion of infection acquired in the	0.32	$(0.47)^{d}$	0.31	(0.46)	0.60	(0.49)	0.42	(0.49)
hospital								
Age at admission	71.6	(16.6)	70.8	(17.0)	55.2	(22.3)	55.3	(21.5)
Length of stay	10.0	(9.5)	9.7	(9.3)	23.4	(21.9)	19.7	(20.0)
Proportion of males	0.60	(0.5)	0.60	(0.49)	0.49	(0.50)	0.49	(0.50)
Proportion of blacks	0.14	(0.35)	0.16	(0.37)	0.21	(0.41)	0.25	(0.44)
Proportion of whites	0.77	(0.42)	0.79	(0.41)	0.66	(0.48)	0.66	(0.48)
Proportion of other race	0.08	(0.28)	0.05	(0.21)	0.13	(0.33)	0.09	(0.29)
Proportion of Medicare / Medicare	0.81	(0.39)	0.81	(0.39)	0.52	(0.50)	0.55	(0.50)
Advantage								
Proportion of Medicaid patients	0.07	(0.25)	0.08	(0.27)	0.16	(0.37)	0.18	(0.39)
Proportion of commercial insurance	0.08	(0.28)	0.07	(0.25)	0.24	(0.43)	0.21	(0.41)
patients								
Proportion other government insurance	0.02	(0.13)	0.02	(0.13)	0.03	(0.16)	0.02	(0.15)
Proportion of self paying patients	0.01	(0.10)	0.01	(0.10)	0.03	(0.16)	0.02	(0.15)
Proportion of other insurances	0.01	(0.11)	0.01	(0.10)	0.03	(0.16)	0.02	(0.13)
Proportion using ICU	0.38	(0.49)	0.40	(0.49)	0.55	(0.50)	0.50	(0.50)
Number of hospitals	2	04	22	25	2	10	22	25

Exhibit 1. Descriptive Statistics of Patients with CAUTI and VCAI Pre-policy (2007Q1–2008Q3) and Post-polic	у
(2008Q4–2011Q3)	

NOTES: a Catheter associated urinary-tract infections

^b Vascular catheter associated infections

^c Number of cases in dataset pre-or-post policy

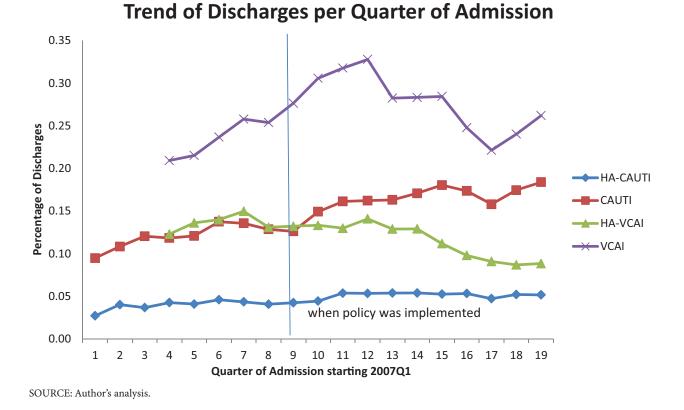
^d Standard Deviation.

SOURCE: Author's analysis.

In Exhibit 2, the rate of CAUTI increased, but the rate of HA-CAUTI was almost the same over the years. The rate of VCAI increased until the end of 2009 and then decreased, but not to pre-policy levels; however, the rate of HA-VCAI post-policy decreased.

In Exhibit 1, the percentage of CAUTI prepolicy was 0.12% and 0.16% post-policy; similarly, the percentage of VCAI pre-policy was 0.24% and 0.30% post-policy, suggesting that both infections increased over time. The proportion of hospitalacquired CAUTI (HA-CAUTI) was 0.32 pre-policy and 0.31 post-policy-suggesting that although infection rates increased, the proportion acquired in the hospital remained the same pre-post policy (Exhibit 3). The proportion of hospitalacquired VCAI (HA-VCAI) was 0.60 pre-policy and 0.45 post-policy, suggesting that although the proportion of discharges with VCAI increased with time, the proportion of VCAI acquired in the hospital decreased (Exhibit 3).

The interrupted time series analysis of the proportion of hospital-acquired CAUTI to all CAUTI discharges (Exhibit 4) confirm (Exhibit 3) that the trend over time is insignificant (0.00036 p-value 0.9158), the change in slope



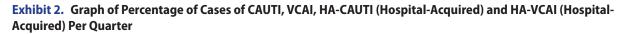
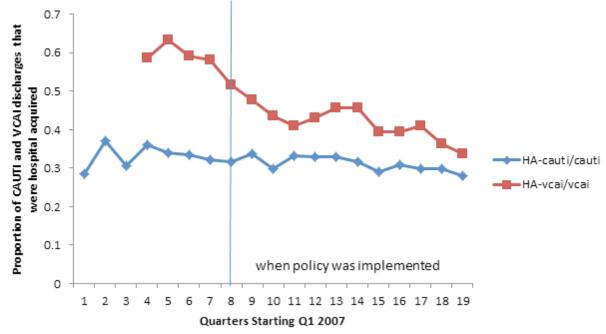


Exhibit 3. Trend of the Proportion of Hospital-Acquired Infections to Discharges of All Respective Infections Per Quarter



SOURCE: Author's analysis.

due to the intervention is insignificant (0.00482 p-value 0.8132), and the slope post-policy (Time*Intervention) is also insignificant (-0.00447 p-value 0.2789). Similarly, the trend over time for analysis of the proportion of hospital-acquired VCAI to all VCAI discharges (Exhibit 4) confirm (Exhibit 3) that there was a significant decline over time (-0.01950 p-value 0.0481), and the change of slope due to the policy (Intervention) was significant (-0.06625 p-value 0.0372). However the change in policy over time (Time*Intervention) was not significant (0.00927 p-value 0.3363), suggesting that the initial decline in rate due to the policy has not been sustained over time.

Exhibit 5 presents the odd ratios from the GHLM logistic regressions of HA-VCAI and HA-CAUTI models.

The HA-VCAI Model

The odds that a patient diagnosed with VCAI acquired in the hospital post-policy was reduced by 43% (OR = 0.57). Other covariates with statistical significance include age (an additional year increases the odds of VCAI by 1.3% OR = 1.013), length of stay (an additional day in the hospital increases your odds by 12% OR = 1.122), and race (Blacks have lower odds compared to Whites, OR = 0.843).

Beneficiaries holding forms of insurance other than Medicare—private insurance (OR = 1.314), other government insurance (OR = 1.33), self pay (OR = 2.058), and other insurance (OR = 1.536) had higher odds than Medicare patients to acquire VCAI in the hospital. Using the ICU increased the odds of patients acquiring VCAI in the hospital by 97% (OR = 1.966).

The HA-CAUTI Model

There was no statistically significant association between the timing of the policy and the odds of acquiring CAUTI in the hospital (OR = 0.968) postpolicy. Results of the other covariates were similar to the VCAI model with some few exceptions: the odds ratio of males to females was 0.548, Medicaid patients were less likely than Medicare patients (OR = 0.824) to acquire CAUTI in the hospital, and there was no difference in the odds of acquiring CAUTI between Medicare patients and other government insurance holders.

The hierarchical model analysis show a level 2 estimate of 0.709 (S.E 0.082) for the VCAI model and 0.705 (S.E 0.103) for the CAUTI model. These estimates are more than twice their standard errors; therefore, some of the variability in the dependent variables is due to differences among hospitals.

Exhibit 4. Interrupted Time 9	Series Analysis of HA	A-CAUTI/CAUTI and H	A-VCAI/VCAI

Variable	Estimate	Standard Error	T-Value	P-Value		
Hospital-acquired CAUTI/CAUTI						
Intercept	0.32800	0.01709	19.19	< 0.0001		
Effect over Time	0.00036	0.00338	0.11	0.9158		
Post-Intervention Effect	-0.00447	0.00398	-1.12	0.2789		
Immediate Intervention Effect	0.00482	0.02004	0.24	0.8132		
Hospital-acquired VCAI/VCAI						
Intercept	0.69860	0.05464	12.79	< 0.0001		
Effect over Time	-0.01950	0.00886	-2.20	0.0481		
Post-Intervention Effect	0.00927	0.00926	1.00	0.3363		
Immediate Intervention Effect	-0.06625	0.02828	-2.34	0.0372		

SOURCE: Author's analysis.

Variable	Description	HA-VCAI Model	HA-CAUTI Model
POLICY	=1 from QTR4 2008	0.571***	0.968
AGE	Age at admission	1.013***	1.010***
LOSDAYS	Length of stay	1.122***	1.143***
GENDER	Proportion of males vs. females	1.052	0.548***
RACE	Proportion of blacks vs. whites	0.831***	0.775***
	Other races vs. white	1.006	0.994
MEDICAID	Medicaid holders vs. Medicare holders	1.045	0.824**
PRIVATE	Private insurance holders vs. Medicare	1.314***	1.942***
	holders		
GOVT	Government insurance holders vs. Medicare	1.330***	1.030
SELFPAY	Self-payers vs. Medicare holders	2.058***	2.083***
OTHINS	Other insurance vs. Medicare holders	1.536***	1.420
ICU	Admitted to ICU vs. not in the ICU	1.966***	2.089***

Exhibit 5. Odd ratios from the logistic regression of the hospital-acquired infections in acute care hospitals in Florida, 2007–2011

NOTES: *** means statistically significant at 1%

** means statistically significant at 5%

SOURCE: Authors' analyses.

Discussion

The goal of this study was to evaluate the association of Medicare non-payment policy on the probability of acquiring certain hospital-acquired infections.

The hypothesis was that the policy would provide a strong enough financial incentive for most hospitals to reduce the probability of patients acquiring one of the infections in the post-policy period. The results of both the interrupted time series and the multivariate regression analysis suggest that there was a reduction in VCAI postpolicy, but not CAUTI. The reduction in VCAI was substantial, suggesting that hospitals in Florida might have responded more strongly to the nonpayment policy for VCAI, which is more expensive to treat than CAUTI (CMS Proposed Rule, 2008). Thus, the perceived magnitude of the financial loss from a hospital-acquired condition may have influenced hospitals' attention and strategy towards infection control. Hartmann et al. (2012), in a

qualitative study on 36 infection preventionists at nonfederal acute care hospitals, reported that there has been increased focus on these hospital-acquired infections and reported some evidence for increased urine analysis at admission since the policy. Central line-associated bloodstream infections (CLABSI), of which VCAI is a part, have received increased attention by not only the federal government, but also other healthcare organizations, such as The Joint Commission (TJC). For instance, the Agency for Healthcare Research and Quality (AHRQ), in collaboration with researchers from Johns Hopkins University, used the comprehensive unit-based safety program (CUSP) to achieve a 40% reduction in CLABSI in 44 states' 1,100 intensive care units over 4 years (AHRQ, 2012), which could have contributed to the decline in VCAI (CUSP was started in 2003 in selected hospitals, but the project was extended to 10 states in 2008 and nationwide in 2009/2010, so it is difficult to predict the likely impact on VCAI independent of the policy). Additionally, hospital infection rate reporting requirements in several states, including Florida, also could have contributed to the observed decline in VCAI post-policy.

It is also possible that medical coders might have contributed to the reduction in cases, especially for VCAI, since VCAI are more expensive to treat. We explored this possibility by looking at the proportion of Medicare patients coded with POA = 0 (i.e., who acquired these infections in the hospital) over time. Since the policy targeted Medicare patients, the rationale is that hospital medical coders will have an incentive to pay increased attention to those patients. Results (not displayed) show that the decrease in the rate of Medicare patients with HA-VCAI was insignificant. Moreover, CMS, in collaboration with RTI and Clarity Coding, conducted research into the accuracy of the POA indicator for hospital-acquired conditions as reported by hospitals. In their report, published in June 2012, the researchers compared hospital MedPar claims and the corresponding medical reports, and found CAUTI POA coding was 94% accurate and VCAI POA coding was 97% accurate (Snow et al., 2012). This adds credence to the use of POA coding for our analysis and the limited role of code manipulation on the results.

The concurrent interventions to reduce hospital-acquired infections, especially in CLABSI, suggest the degree of association we found could be overestimated. The findings that VCAI appeared to be more responsive to the policy than CAUTI, however, suggest that the impact of the financial incentive, if any, may depend on its magnitude, as VCAI is relatively more expensive to treat. Thus, the specificity of the financial incentive may have contributed (Young, 2005); that is, Medicare was specific about the type of conditions covered by the policy and hospitals knew how much noncompliance with the policy would cost.

Although the financial incentive might have affected different conditions differently and could be overestimated, because of concurrent interventions, the initial response is encouraging. And, as Medicare expands the non-payment policy to more hospital-acquired conditions, hospitals are more likely to continue to make changes that will positively affect patient outcomes.

It is important to note some limitations of this study. The most important limitation is the lack of a non-policy control group that could have allowed a difference-in-difference analysis. It is possible that infection rates were declining over this period independent of the Medicare policy, but because there was no suitable control group of acute care hospitals that were not affected by this policy, it is not possible to determine this. Secondly, we were limited in the duration of prepolicy data, because there were no POA data prior to 2007 in Florida, and VCAI was a new requirement that was reported starting in the fourth quarter of 2007. Additionally, because this is an observational study, it is subject to all the limitations of non-randomized studies and, therefore, causality cannot be implied.

Furthermore, the study was limited to Florida hospitals, nonetheless, it is one of the largest states in the U.S. and the data included a wide range of hospital types. We were unable to identify a national dataset that included information on whether infections were POA during the pre-policy period. In addition, the validity of the study is predicated on the accuracy of the DRG and the POA coding (Zhan et al., 2009). Some researchers have also expressed concern about adequacy of medical coding for research purposes, because, by law, medical coders can only code from information in the medical record documented by a provider (physicians, physician assistants, and nurse practitioners), which means nursing notes that might include information related to hospital-acquired infection will not be coded (Meddings et al., 2012). Because VCAI was a new requirement, there were no cases in the first 3 quarters of 2007 and there could have

been a learning curve for medical personnel in reporting cases. It is not known what impact that could have had on the results. However, even with these limitations, this study demonstrates that actual (as opposed to simulated) VCAI post-policy infection rates significantly declined in Florida. Our results are similar to what Lee *et al.* (2012) found: a significant reduction in CLABSI, but they attributed the decline to other factors, because CLABSI was also declining before the policy was implemented.

Additional research is warranted to further understand the underlying factors that led to the significant improvement in VCAI and the significant increase in rates for private insurance patients. In the future, when Medicare releases post-policy data at the national level, it will be possible to explore the organizational changes made by hospitals in order to better understand behavioral changes due to the policy.

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References

- Agency for Health Care Administration (2010). Retrieved from http://www.floridahealthfinder. gov/Researchers/OrderData/order-data.aspx
- Agency for Healthcare Research and Quality (2012). Retrieved from http://www.ahrq.gov/ news/press/pr2012/pspclabsipr.htm
- Apisarnthanarak, A., Thongphubeth, K.,
 Sirinvaravong, S., Kitkangvan, D., Yuekyen,
 C., Warachan, B., . . . Fraser, V. J. (2007). The
 Effectiveness of Multifaceted Hospital Wide
 Quality Improvement Programs Featuring an
 Intervention to Remove Unnecessary Urinary

Catheters at a Tertiary Care Center in Thailand. Infection Control and Hospital Epidemiology, 28(7), 791–798. PubMed http://dx.doi. org/10.1086/518453

- Burke, J. P. (2003). Infection control-A problem for patient safety. *The New England Journal of Medicine*, 348, 651–656. PubMed http://dx.doi. org/10.1056/NEJMhpr020557
- Centers for Medicare & Medicaid Services (2006). Eliminating Serious, Preventable, and Costly Medical Errors-Never Events. Retrieved from http://www.cms.hhs.gov/apps/media/press/ release.asp?counter=1863
- Centers for Medicare & Medicaid Services (2008). Roadmap for Implementing Value Driven Healthcare in the Traditional Medicare Fee-for-Service Program. Retrieved from http://www.cms.gov/Medicare/Quality-Initiatives-Patient-Assessment-Instruments/ QualityInitiativesGenInfo/downloads/ vbproadmap_oea_1-16_508.pdf
- ECRI Institute Special Advisory (2008). List of CMS Hospital-Acquired Conditions Expanded under New Final Rule. Retrieved from https:// www.ecri.org/Documents/Patient_Safety_ Center/CMS_New_Final_Rule.pdf
- CMS Proposed Rules, 73 Fed. Reg. 23528–23938 (proposed Apr. 30, 2008) (to be codified at 42 C.F.R. pt. 411, 412, 413, 422, 489. Retrieved from http://edocket.access.gpo.gov/2008/pdf/ 08-1135.pdf
- Greene, M. T., Chang, R., Kuhn, L., Rogers, M. A., Chenoweth, C. E., Shuman, E., & Saint, S. (2012).
 Predictors of Hospital-Acquired Urinary Tract-Related Bloodstream Infection. *Infection Control and Hospital Epidemiology*, 33(10), 1001–1007.
 PubMed http://dx.doi.org/10.1086/667731

MMRR

- Haley, R. W., Quade, D., Freeman, H. E., & Bennett,
 J. V. (1980). Study on the Efficacy of Nosocomial Infection Control (SCENIC PROJECT):
 Summary of Study Design. *American Journal of Epidemiology*, 111(5), 472–485. PubMed
- Hartmann, C. W., Hoff, T., Palmer, J. A., Wroe, P., Dutta-Linn, M. M., & Lee, G. (2012). The Medicare Policy of Payment Adjustment for Health Care Associated Infections: Perspectives on Potential Unintended Consequences. *Medical Care Research and Review*, 69(1), 45–61. PubMed http://dx.doi.org/10.1177/1077558711413606
- Institute for Health Metrics (2009). CMS Reimbursement Rules for Hospital-Acquired Infections: How Electronic Reports Can Protect the Bottom Line. Retrieved from http://www. lifesciencewriter.com/Getting_Ahead_of_ Hospital_Acquired_Infection.pdf
- Johnson, J. R., Kuskowski, M. A., & Wilt, T. J. (2006).
 Systematic Review: Antimicrobial Urinary Catheters to Prevent Catheter-Associated Urinary Tract Infection in Hospitalized Patients.
 Annals of Internal Medicine, 144(2), 116–126.
 PubMed http://dx.doi.org/10.7326/0003-4819-144-2-200601170-00009
- Kaye, K. S., Marchaim, D., Chen, T. Y., Chopra, T., Anderson, D. J., Choi, Y., & Schmader, K. E. (2011). Predictors of Nosocomial Bloodstream Infections in Older Adults. *Journal of the American Geriatrics Society*, 59(4), 622–627. PubMed http:// dx.doi.org/10.1111/j.1532-5415.2010.03289.x
- Lee, G. M., Kleinman, K., Soumerai, S. B., Tse, A., Cole, D., Fridkin, S. K., . . . Jha, A. K. (2012). Effect of Nonpayment for Preventable Infections in U.S. Hospitals. *The New England Journal of Medicine*, *367*, 1428–1437. PubMed http:// dx.doi.org/10.1056/NEJMsa1202419

- McHugh, M., Martin, T. C., Orwat, J., & Van Dyke,
 K. (2011). Medicare's policy to Limit Payment for
 Hospital-Acquired Conditions: Impact on Safety
 Net Providers. *Journal of Health Care for the Poor and Underserved*, 22(2), 638–647. PubMed
 http://dx.doi.org/10.1353/hpu.2011.0058
- McNair, P. D., Luft, H. S., & Bindman, A. B. (2009). Medicare's Policy Not To Pay For Treating Hospital-Acquired Conditions: The Impact. *Health Affairs*, 28(5), 1485–1493. PubMed http://dx.doi.org/10.1377/hlthaff.28.5.1485
- McNutt, R., Johnson, T. J., Odwazny, R., Remmich,
 Z., Skarupski, K., Meurer, S., . . . Harting, B.
 (2010). Change in MS-DRG Assignment and
 Hospital Reimbursement as a Result of Centers for Medicare & Medicaid Changes in Payment for Hospital-Acquired Conditions: Is it Coding or Quality? *Quality Management in Health Care*, 19(1), 17–24. PubMed
- Meddings, J., Saint, S., & McMahon, L., Jr. (2010, June). Hospital-Acquired Catheter-Associated Urinary Tract Infection: Documentation and Coding Issues May Reduce Financial Impact of Medicare's New Payment Policy. *Infection Control and Hospital Epidemiology*, 31(6), 627– 633. PubMed http://dx.doi.org/10.1086/652523
- Meddings, J. A., Reichert, H., Rogers, M. A., Saint,
 S., Stephansky, J., & McMahon, L. F. (2012).
 Effect of Non-Payment for Hospital-Acquired,
 Catheter-Associated Urinary Tract Infection: A
 Statewide Analysis. *Annals of Internal Medicine*,
 157(5), 305–312. PubMed http://dx.doi.org/
 10.7326/0003-4819-157-5-201209040-00003
- Pacemakerclub. (2012). Medicare Ruling Concerns Hospital CEOs. Retrieved from http://www. pacemakerclub.com/public/jpage/l/p/story/a/ Storypage/sid/25334/content.do

- Petersen, L. A., Woodard, L. D., Urech, T., Daw,
 C., & Sookanan, S. (2006). Does Pay-forPerformance Improve the Quality of Healthcare?
 Annals of Internal Medicine, 145(4), 265–272.
 PubMed http://dx.doi.org/10.7326/0003-4819145-4-200608150-00006
- Pronovost, P., Needham, D., Berenholtz, S., Sinopoli, D., Chu, H., Cosgrove, S., ... Goeschel, C. (2006). An Intervention to Decrease Catheter-Related Bloodstream Infections in the ICU. *The New England Journal of Medicine*, 355(26), 2725–2732. PubMed http://dx.doi.org/10.1056/ NEJMoa061115
- Raudenbush, S., & Bryk, A. (2002). Hierarchical Linear Models: Applications and Data Analysis Methods, London: Sage Publications.
- Scott, R. D. (2002). The Direct Medical Cost of Healthcare-Associated Infections in U.S. Hospitals and the Benefits of Prevention.CDC Division of Healthcare Quality Promotion. Retrieved from http://www.cdc.gov/ncidod/ dhqp/pdf/scott_costpaper.pdf
- Snow, C. L., Holtzman, L., Waters, H., McCall, N. T., Halpern, M., & Newman, L. ... Guzman, C. R. (2012). Accuracy of Coding in the Hospital-Acquired Conditions Present on Admission ProgramFinalReport.Retrievedfromhttp://www. cms.gov/Medicare/Medicare-Fee-for-Service-Payment/HospitalAcqCond/Downloads/ Accuracy-of-Coding-Final-Report.pdf

- Stone, P. W., Gilead, S., McNair, P., Mattes, N., Cohen, B., Landers, T., & Larson, E. (2010).
 CMS Changes in Reimbursement for HAIs: Setting a Research Agenda. *Medical Care*, 48(5), 433–439. PubMed http://dx.doi.org/10.1097/ MLR.0b013e3181d5fb3f
- Wagner, A. K., Soumarai, S. B., Zhang, F., & Ross-Degnan, D. (2002). Segmented Regression Analysis of Interrupted Time Series Studies in Medication Use Research. *Journal of Clinical Pharmacy and Therapeutics*, 27(4), 299–309.
 PubMed http://dx.doi.org/10.1046/j.1365-2710.2002.00430.x
- Wald, H. L., & Kramer, A. M. (2007). Nonpayment for Harms Resulting from Medical Care. *Journal* of the American Medical Association, 298(23), 2782–2784. PubMed http://dx.doi.org/10.1001/ jama.298.23.2782
- Young, G. J. (2005). Conceptual Issues in the Design and Implementation of Pay-for-Quality Programs. *American Journal of Medical Quality*, 20(3), 144–150. PubMed http://dx.doi. org/10.1177/1062860605275222
- Zhan, C., Elixhauser, A., Richards, C. L., Jr., Wang, Y., Baine, W. B., Pineau, M., . . . Hunt, D. (2009). Identification of Hospital-Acquired Catheter-Associated Urinary Tract Infections from Medicare Claims. *Medical Care*, 47(3), 364–369. PubMed http://dx.doi.org/10.1097/ MLR.0b013e31818af83d